

# Short-term earthquake probabilities during the L'Aquila earthquake sequence in central Italy, 2009

### S23A-4479

### S23A: Aftershock Hazard: Forecasting Highly Time- and Space-Dependent Seismicity and Shaking after Large Mainshocks

# Introduction

We apply two versions of the ETAS model and the Proximity to Past Earthquakes (PPE) model as the reference model to the Italian Seismic Instrumental and Parametric Database (ISIDE) of shallow seismicity collected by INGV for the 2009 L'Aquila seismic sequence that occurred in Abruzzi Region (Central Italy) (Figure 1) for the purpose of understanding the role of small events. The data set spans from March 16, 2009 to June 30, 2009. We use the probability gain to evaluate their forecasting performance.

L'Aquila foreshock sequence, starting at the end of 2008 (with the largest event, Mw4.0, on March 30, 2009), was characterized by clustering around its mainshock (April 6, 2009, M<sub>w</sub>6.3) nucleation area. Two other strong events with their own aftershocks occurred after the mainshock, on April 7, 2009 ( $M_w$  5.6) and April 9, 2009 ( $M_w$  5.4), respectively.

## Models for earthquake occurrences

## PPE (Proximity to Past Earthquakes) model

The PPE model is a Poisson model with a specific method of estimating the seismicity rate, i.e., it is a smoothed seismicity baseline model. It can play the role of a spatially varying reference model against which the performance of time-varying models can be compared. For this reason, it is adopted in this study as the reference model. It was proposed/formulated by Jackson and Kagan (1999) and named by Rhoades and Evison (2004).

$$\lambda(t, x, y, m) = \frac{s(m)}{t - t_0} \sum_{t_i < t} \left( \frac{a}{d^2 + (x - x_i)^2 + (y - y_i)^2} + \epsilon \right), \ m \ge m_c$$

 $s(m) = \beta e^{-\beta(m-m_c)}, \ m \ge m_c$ 

## Space-time ETAS models

Two versions of the ETAS model, named ETAS I and ETAS II, developed by different research groups, were used for evaluating the influence of small earthquakes on predictive performance. The difference between ETAS I and ETAS II is as follows

ETAS I was developed by Ogata (1998) and Console et al. (2010a,b).

ETAS II was developed by Zhuang et al. (2005) and Ogata and Zhuang (2006), and is an improved version of the model in Ogata (1998).

Although these two versions belong to the same class of ETAS model, their details, such as how to determine background seismicity, how to deal with boundary conditions, and how to estimate parameters, are quite different.

Three versions of ETAS II are considered in order to understand the role of small events in forecasting large events, and to verify the statement made by Helmstetter (2003) and Helmstetter et al. (2005) that small events are important in triggering large earthquakes.

ETAS II-2.0+ (events of  $\geq$  M2.0)

ETAS II-1.6+ (events of  $\geq$  M1.6)

ETAS II-2.0m (events of M>1.6 for calculating the expected number of events before t and auxiliary events with 1.6 < M<2.0 for computing the likelihood).

We consider two forms for the spatial probability density function (p.d.f.):

ETAS I.

 $f(v, w; m) = \frac{q-1}{\pi D e^{\alpha(m-m_c)}} \left[ 1 + \frac{v^2 + w^2}{D e^{\alpha(m-m_c)}} \right]^{-q},$ 

 $f(v,w;m) = \frac{q-1}{\pi D e^{\gamma(m-m_c)}} \left[ 1 + \frac{v^2 + w^2}{D e^{\gamma(m-m_c)}} \right]^{-q}.$ 

The parameters to be estimated in the learning phase by the maximum likelihood methods are: (A, $\alpha$ , c, p, D, q) for ETAS I and (A,  $\alpha$ , c, p, D, q,  $\gamma$ ) for ETAS II. In ETAS I, the scale factor for the aftershock productivity function is the same as the scaling factor for the spatial response function.

For ETAS I we assume  $\alpha = 1.0$ ·ln 10 for the physical meaning that the number of triggered events is proportional to the rupture area of the triggering earthquake (see Console et al., 2006; Hainzl et al., 2008).

- $\rightarrow$  Expected number of direct offspring produced by an event of the threshold magnitude  $\alpha \rightarrow$  quantifies the difference in the productive efficiency from events of different magnitude
- c and  $p \rightarrow parameters$  in the Omori-Utsu formula for aftershock frequencies

 $D \rightarrow$  characteristic triggering distance

 $q \rightarrow$  spatial decay coefficient

 $\gamma \rightarrow$  spatial scaling factor

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## Scoring procedures

Given forecasts from different models, it is important to know which model performs the best in forecasting. We only consider the information score, also called the entropy score to evaluate the performance of probability forecasts

The whole space-time-magnitude window for a forecast is divided into cells of equal size. The forecast gives a probability pi that at least one event occurs in the ith space-time-magnitude cell; the reference model, is the Poisson model, which gives a probability of  $\overline{p}_i$ 

The binomial information score is defined as the logarithm of the likelihood ratio of the forecasting and reference

$$B_i^{(b)} = Y_i \log(p_i/\bar{p}_i) + (1 - Y_i) \log \frac{(1 - p_i)}{(1 - \bar{p}_i)}$$

The Poisson information score is defined

$$B_i^{(P)} = \sum_k I(n_i = k) \log(p_{i,k}/\bar{p}_{i,k}),$$

The information gain per unit space-time-magnitude volume is defined as:

 $G = \frac{\Delta}{2} \sum r$ 

V is the total volume of the space-time-magnitude range,  $B_i$  takes on the value of  $B_i^{(b)}$  or  $B_i^{(p)}$ , and  $\Delta$  is the size of each cell. When B<sub>i</sub> takes on the value of B<sup>(b)</sup>, G is called the binomial information gain. Similarly, G is the Poisson information gain when we use  $B_i^{(p)}$ .

	Binomial information scores									
	Model	M2.0+			M3.0+			M4.0+		
	MOUEI	$\mathrm{EC}$	NEC	Tot.	$\mathrm{EC}$	NEC	Tot.	$\mathrm{EC}$	NEC	Tot.
	PPE1.6+	0.83	0.31	1.14	2.96	-0.23	2.72	1.17	-0.06	1.11
	ETAS I	16.15	-5.78	10.37	71.17	-22.21	48.96	35.90	-7.50	28.39
	ETAS II- $1.6+$	18.12	-4.50	13.62	66.63	-17.27	49.36	31.00	-5.25	25.75
	ETAS II- $2.0+$	14.09	-1.09	13.01	56.02	-9.35	46.67	30.11	-3.76	26.35
	ETAS II-2.0m	17.87	-3.43	14.44	64.81	-14.71	50.10	32.13	-5.08	27.05
	Poisson information scores									
	Model	M2.0+			M3.0+			M4.0+		
		EC	NEC	Tot.	$\mathrm{EC}$	NEC	Tot.	$\mathrm{EC}$	NEC	Tot.
	PPE1.6+	215.79	0.31	216.10	12.82	-0.23	12.59	1.67	-0.06	1.60
	ETAS I	7992.22	-5.78	7986.44	666.64	-22.21	644.44	95.18	-7.50	87.68
	ETAS II- $1.6+$	6860.19	-4.50	6855.69	549.77	-17.27	532.50	75.67	-5.24	70.42
	ETAS II- $2.0+$	7322.93	-1.08	7321.84	593.74	-9.35	584.39	81.88	-3.76	78.11
	ETAS II-2.0m	7541.82	-3.43	7541.82	617.48	-14.71	602.77	85.19	-5.08	80.11

**Table 3.** Information gain compared to Model PPE2.0+ for each model. The boldface fonts show the best models in the corresponding category. EC stands for "event cell" and NEC stands for "non-event cel"

#### onclusions Discussion and

All the ETAS models used in this study underestimate the occurrence rate soon after the mainshock and during the aftershock sequence (Figure 3).

Two possible causes could be:

1) the ETAS models were fit to a learning data set of relatively moderate seismic activity, since there are no earthquakes of magnitude 5.0+ in the catalog before the test phase, March 15, 2009. In this case, our maximum likelihood (ML) parameters, applied to larger magnitude mainshocks, are in some sense the results of an extrapolation, which can easily cause large biases

2) the ML parameters were not updated at the end of each day, just before the computation of a new forecast for the following day. This would improve the performance of an ETAS model, especially in the case of a long aftershock sequence

Comparing the two types of ETAS models: 1) ETAS I (with the same fixed exponent coefficient  $\alpha$  = 2.3 for both the productivity function and the scaling factor in the spatial response function) performs better in forecasting the active aftershock sequence than the ETAS II-type models, when the Poisson score is adopted; 2) Before the mainshock all ETAS II models performed better than ETAS I.

The occurrence of foreshock activity some days before the mainshock raised the probability gain up to an order of 20, using the ETAS I model. This does not appear to be a great achievement towards the potential use of this information for civil protection purposes.

We do not consider the performance of these models in forecasting spatial locations, since the area is quite small, 1.1° by 1.5°, while the location error is several kilometers. Moreover, all the models use the G-R magnitude-frequency relation for the magnitude distribution, independent of the time and location components. Thus, there are no differences between these Etas models in forecasting the magnitude distribution of future events.

In summary, to achieve better forecasts, as well-known from the existing literature on the ETAS model, the training catalog should contain some typical examples of magnitudes similar to that of the mainshock. If not, it would be better to use a typical set of parameters from a nearby region or from a region of similar tectonic structure or equivalent seismicity levels. Other lessons from this study are as follows. (1) Making use of the earthquakes of lower magnitude works better than using only events larger than the target magnitude threshold. (2) Parameters should be updated during the forecast process, possible. (3) Forecasts should be updated just after the occurrence of potential foreshock activity, whenever significant moderate earthquakes are observed.

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